

AMBLYOPIA TREATMENT STUDY (ATS23)

A Randomized Trial of Dichoptic Treatment for Amblyopia in Children 4 to 7 Years of Age

Statistical Analysis Plan

Version 1.0

Version History

| Version | Protocol Version | Author | Approver | Effective Date | Study Stage |
|---------|------------------|--------------------|-----------------|----------------|----------------------------------------------|
| 1.0 | 1.0 | Desirae Sutherland | Wesley Beaulieu | 4/17/2024 | First participant has not yet been enrolled. |

| Version | Revision Description |
|---------|----------------------|
| 1.0 | Original Version |

Approvals

| Role | Digital Signature or Handwritten Signature/Date |
|---------------------------------------------------------------------|-------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------|
| Author: Desirae Sutherland | <p style="text-align: center;">Desirae Sutherland</p> <small>Digitally signed by Desirae Sutherland DN: cn=Desirae Sutherland ou=North Wing Reason: I am the author of this document Location: Date: 2024-04-18 14:10-04:00</small> |
| Senior Statistician: Wesley Beaulieu | <p style="text-align: center;">Wesley Beaulieu</p> <p style="text-align: center;">I agree to the terms defined by the placement of my signature in this document</p> <p style="text-align: center;">2024-04-18 09:55-04:00</p> |
| Jaeb Principal Investigator: Raymond Kraker | <p style="text-align: center;">Ray Kraker</p> <p style="text-align: center;">I agree to the terms defined by the placement of my signature in this document</p> <p style="text-align: center;">2024-04-18 10:13-04:00</p> |
| Study Principal Investigator: Robert Henderson | <p style="text-align: center;">Robert Henderson</p> <small>Digitally signed by Robert Henderson DN: cn=Robert Henderson ou=South Wing Reason: I have reviewed this document Location: Date: 2024-04-18 10:18-04:00</small> |

1 **1. Study Overview**

2 This document outlines the statistical analyses to be performed for the ATS23 Trial and to be
3 included in the primary manuscript data packet.

4 The protocol is a multicenter trial designed to compare the change in amblyopic eye distance VA
5 from randomization to 26 weeks in participants randomized to treatment with the Luminopia
6 One headset (1 hour per day of watching dichoptic movies, 6 days per week) or patching (2
7 hours of patching per day, 7 days per week).

8 The aforementioned treatment regimens of Luminopia One and patching shall subsequently be
9 referred to as LUMINOPIA and PATCHING, respectively.

10 Approximately 238 participants will be enrolled and randomized in a 1:1 ratio to either
11 LUMINOPIA or PATCHING treatment (119 per group).

12 At the 26-week primary outcome visit, participants who were randomly assigned to receive
13 PATCHING treatment and have an interocular difference (IOD) of 1 line (0.10 logMAR) or
14 more will be offered Luminopia dichoptic therapy. Participants who agree to the treatment will
15 continue follow-up with visits at 39- and 52-weeks post-randomization. Otherwise, for all other
16 participants the study will end at 26 weeks.

17 The study protocol specifies the eligibility criteria and schedule of study visits and procedures.

18 **2. Statistical Hypotheses**

19 The primary efficacy outcome will be the change in amblyopic eye distance VA (measured as
20 logMAR) from randomization to 26 weeks. Change in logMAR will be calculated as [outcome
21 VA] – [randomization VA] such that a negative change indicates improvement in visual acuity,
22 and a positive change indicates worsening.

23 The study is designed to test a one-sided null hypothesis that the mean change in VA from
24 baseline at 26 weeks with LUMINOPIA is inferior to PATCHING by 0.0625 logMAR (i.e., 5/8
25 of one line) or more in favor of the alternative hypothesis that LUMINOPIA is non-inferior to
26 PATCHING.

27 **$H_0: \mu_{\text{PATCHING}} - \mu_{\text{LUMINOPIA}} \leq -0.0625 \text{ logMAR}$** (LUMINOPIA inferior to PATCHING)

28 **$H_a: \mu_{\text{PATCHING}} - \mu_{\text{LUMINOPIA}} > -0.0625 \text{ logMAR}$** (LUMINOPIA *not* inferior to PATCHING)

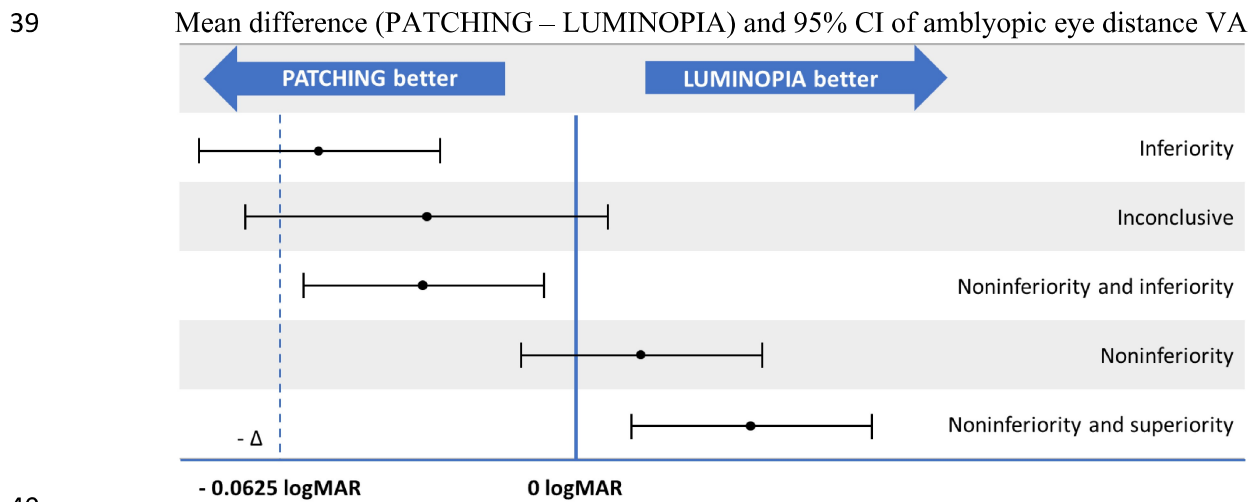
29 To represent the difference between treatment groups (PATCHING minus LUMINOPIA), a two-
30 sided 95% confidence interval (CI) will be constructed. Since the LOWER limit of a two-sided

31 95% CI is equivalent to the LOWER limit of a one-sided 97.5% CI, this will allocate a
32 significance level of 0.025 to be used in testing noninferiority.

33 Non-inferiority of LUMINOPIA to PATCHING will be declared if the LOWER limit of the two-
34 sided 95% CI for the difference between treatment groups is greater than the non-inferiority limit
35 of -0.0625 logMAR favoring PATCHING (Figure 1).

36 If non-inferiority is declared, a test of no difference (superiority test) in mean VA change at 26
37 weeks for LUMINOPIA compared to PATCHING will be conducted.

38 **Figure 1. Depiction of Null and Alternative Hypotheses for Treatment Group Difference**



41 3. Outcome Measures

42 3.1. Primary Efficacy Endpoint

- 43 • Change in amblyopic eye distance VA from baseline at 26 weeks

44 3.2. Secondary Efficacy Endpoints

- 45 • Change in child and proxy PedEyeQ Functional Vision domain scores from baseline at
46 26 weeks
- 47 • Change in child and proxy PedEyeQ Social domain scores from baseline at 13 weeks
- 48 • Change in child and proxy PedEyeQ Frustration/Worry domain scores from baseline at
49 13 weeks

50 3.3. Exploratory Efficacy Endpoints

- 51 • Change in binocular function score from baseline at 13 weeks and 26 weeks
- 52 • Change in amblyopic eye distance VA from baseline at 13 weeks

- 53 • Change in amblyopic eye distance VA from baseline over 26 weeks (area under the
54 curve)
- 55 • Improvement of amblyopic eye distance VA by 2 or more lines (0.2 logMAR) from
56 baseline at 13 and 26 weeks
- 57 • Resolution of amblyopia from baseline at 13 and 26 weeks
- 58 • Child, proxy, and parent Treatment Impact Questionnaire scores at 13 and 26 weeks

59 Post-primary Outcome Follow Up

- 60 • Change in binocular function score from 26 weeks at 39 weeks and 52 weeks
- 61 • Change in amblyopic eye distance VA from 26 weeks at 39 weeks and 52 weeks
- 62 • Improvement of amblyopic eye distance VA by 2 or more lines (0.2 logMAR) from 26
63 weeks at 39 and 52 weeks
- 64 • Resolution of amblyopia from 26 weeks at 39 and 52 weeks

65 **4. Description of Statistical Methods**

66 Analyses will follow the intent-to-treat principle (ITT); all participants will be analyzed
67 according to their randomized treatment group, irrespective of adherence or compliance.
68 However, a per protocol analysis will be performed for the primary outcome to check sensitivity
69 of the results (section 5.2.1). The intent-to-treat analysis is considered primary and if the results
70 of the per-protocol analysis and intent-to-treat give inconsistent results, exploratory analyses will
71 be performed to evaluate possible factors contributing to the differences.

72 **5. Primary Efficacy Outcome**

73 **5.1. Analysis of the Primary Endpoint**

74 The primary endpoint, change in amblyopic eye logMAR distance VA from baseline at 26
75 weeks, is a continuous variable that will be analyzed using an analysis of covariance (ANCOVA)
76 model to estimate the adjusted mean difference between PATCHING and LUMINOPIA. The
77 model will adjust for baseline amblyopic-eye distance VA and prior treatment for amblyopia
78 (glasses only vs other treatment in addition to glasses). The adjusted between-group mean
79 difference, two-sided 95% CI and *p*-value for a test of no difference will be reported.

80 Participants who do not complete the 26-week visit will have their 26-week amblyopic eye
81 distance VA imputed. Markov chain Monte Carlo multiple imputation with 100 imputations will
82 be used to impute missing data; variables in the imputation model will include prior treatment for
83 amblyopia and amblyopic-eye VA at baseline, 13, and 26 weeks. Imputation will be carried out
84 separately for PATCHING and LUMINOPIA. Reasons for which a participant may not complete
85 the 26-week visit are outlined in section 8, “Primary Estimand.”

86 Non-inferiority of LUMINOPIA compared to PATCHING will be declared if the LOWER limit
87 of the two-sided 95% CI for the difference between treatment groups (PATCHING minus
88 LUMINOPIA) in mean change in logMAR distance VA from baseline to 26 weeks is greater
89 than (i.e., to the right of) the non-inferiority limit of -0.0625 logMAR favoring PATCHING.
90 Note that the LOWER limit of a two-sided 95% confidence interval is equivalent to the lower
91 limit of a one-sided 97.5% confidence interval.

92 If non-inferiority is demonstrated, superiority of LUMINOPIA over PATCHING will be
93 declared if the LOWER limit of the 95% CI is greater than zero. Conversely, superiority of
94 PATCHING over LUMINOPIA will be declared if the UPPER limit of the 95% CI is less than
95 zero; in this scenario, LUMINOPIA will be both inferior to PATCHING and non-inferior to
96 PATCHING with a margin of 0.0625 logMAR distance VA (Figure 1).

97 A boxplot showing changes in VA at 13 and 26 weeks by treatment group will be presented to
98 aid in interpretation.

99 **5.2. Sensitivity Analyses of the Primary Endpoint**

100 To explore the robustness of the primary analysis, sensitivity analyses will be conducted and are
101 outlined below. All sensitivity analyses will control for baseline amblyopic-eye VA and prior
102 treatment for amblyopia.

103 **5.2.1. Complete cases (Sensitivity Analysis #1)**

104 The primary outcome will be analyzed without imputation of missing data.

105 **5.2.2. Per protocol (Sensitivity Analysis #2)**

106 The primary outcome will be analyzed using the same methods but participants who
107 discontinued their assigned treatment or initiated non-randomized treatment will be excluded
108 from both the imputation and analysis models.

109 **5.2.3. Outliers (Sensitivity Analysis #3)**

110 To ensure that statistical outliers do not have undue impact on analyses, the change in distance
111 VA from baseline at 26 weeks will be modeled with robust regression using the Huber M-
112 estimator instead of ANCOVA. Missing data will still be imputed using multiple imputation.

113 **5.2.4. Confounding (Sensitivity Analysis #4)**

114 To ensure that confounding does not affect study results, if an imbalance of baseline factors
115 between treatment groups is observed, the primary analysis will be repeated, controlling for these
116 potential confounders. The determination of a meaningful baseline imbalance will be based on
117 clinical judgement.

118 **5.2.5. Heteroscedasticity (Sensitivity Analysis #5)**

119 To ensure that heteroscedasticity does not affect study results, a linear model that applies the
120 residual-based estimator HC3 (Firores) will be used to estimate the empirical standard error.
121 Note that this will be used as an alternative to model-based standard error which may be
122 incorrect in the case of severe heteroscedasticity. The model will produce the adjusted between-
123 group mean difference of the change in VA at 26-weeks and two-sided 95% CI. Missing data
124 will still be imputed using multiple imputation.

125 **6. Secondary Efficacy Outcomes**

126 Secondary analyses will test the null hypothesis of no difference between treatment groups.

127 **6.1. Pediatric Eye Disease Questionnaire (PedEyeQ)**

128 Quality of life will be evaluated for children respondents aged 5 to 7 years in each treatment
129 group using the PedEyeQ questionnaire. Additionally, the parent will answer on behalf of his/her
130 child as a proxy for children 4 to 7 years of age. Scores on Functional Vision, Frustration/Worry,
131 and Social domains will be assessed for both child and proxy at baseline as well as at the visit
132 week indicated below (Table 2). Responses will be Rasch scored according to reference tables
133 and standardized on a ratio scale ranging from 0 to 100.¹

134 **Table 2. Structure of the PedEyeQ Analysis: Domains and Respondents**

| Participant Age | Respondent Level | Domain | | | Outcomes |
|-----------------|------------------|-------------------|------------------------------|------------------------------|----------|
| | | Social (13 weeks) | Frustration/Worry (13 weeks) | Functional Vision (26 weeks) | |
| 4-7 years | Proxy | 1 | 1 | 1 | 3 |
| 5-7 years | Child | 1 | 1 | 1 | 3 |

Total = 6

135 Univariate analysis of covariance (ANCOVA) will be used to assess the difference between
136 treatment groups across all domains and respondents (3 domains × 2 respondents = 6 outcomes)
137 as shown in Table 2. Models will be adjusted for prior treatment for amblyopia and enrollment
138 scores. The treatment effect will be summarized as a mean difference and 95% CI.

139 PedEyeQ scores will be imputed for any participants who did not respond to the questionnaires
140 at the 13-week or 26-week visits. Markov chain Monte Carlo multiple imputation with 100
141 imputations will be used to impute missing scores for each domain. The imputation model will
142 include 12 variables to represent the three domains (Functional Vision, Frustration/Worry, and
143 Social) and two levels of respondents (proxy and child) at the enrollment and outcome visits, and
144 will include a variable to indicate prior treatment for amblyopia. Imputation will be carried out
145 separately for PATCHING and LUMINOPIA.

146 **7. Visit Windows**

147 To be included in analyses, visits must be completed within the specified visit windows (Table
 148 3). Values from visits outside of the indicated analysis window will be considered missing data.

149 **Table 3. Analysis Windows for Primary and Post-primary Outcome Follow-up Visits**

| Primary Outcome Follow-up (Randomization to 26-Week Visit) | | |
|------------------------------------------------------------------------|------------------------------------------|-------------------------------------------------------------------|
| Visit | Target Day Post-Randomization | Analysis Window around Target Day |
| 13-Week Visit | 91 days | ± 5 weeks (56 days to 125 days) (8 weeks to <18 weeks) |
| 26-Week Visit Primary Outcome | 182 days | ± 8 weeks (126 days to 238 days) (18 weeks to 34 weeks) |
| Post-primary Outcome Follow-up (39-Week and 52-Week Visits) | | |
| Visit | Target Day Post 26-week Visit | Analysis Window around Target Day |
| 39-Week Visit | 91 days | ± 5 weeks (56 days to 125 days) (8 weeks to <18 weeks) |
| 52-Week Visit | 182 days | ± 8 weeks (126 days to 238 days) (18 weeks to 34 weeks) |

150 **8. Primary Estimand**

151 The primary outcome is derived from VA measurements at 26 weeks. The clinical question is
 152 whether the change in VA with LUMINOPIA is non-inferior to PATCHING with a non-
 153 inferiority margin of 0.0625 logMAR. The population-level summary measure is the mean
 154 difference comparing the LUMINOPIA and PATCHING groups.

155 Table 4 specifies the foreseen intercurrent events, whether data will be imputed after the event,
 156 and the strategy as defined in *E9(R1) Statistical Principles for Clinical Trials: Addendum:
 157 Estimands and Sensitivity Analysis in Clinical Trials*. Data that are missing due to death, loss to
 158 follow-up, or participant withdrawal will be imputed based on observed VA measurements. This
 159 is consistent with a hypothetical scenario in which the intercurrent events do not occur and
 160 assumes that outcomes in those dying, lost to follow-up, and withdrawn resemble outcomes of
 161 those without missing data due to these events. Treatment discontinuation, treatment crossover,
 162 receipt of treatment for a condition other than amblyopia, and receipt of alternative treatment for
 163 amblyopia allow for continued observation of the outcome but might affect the outcome itself.
 164 By using observed data from participants who experience these events, we are adopting a

165 treatment policy strategy in which the value for the variable of interest is used regardless of
 166 whether the intercurrent event occurs. This strategy aligns with the ITT Principle.

167 **Table 4. Intercurrent Events, Censoring, and Treatment Effects for the Primary Outcome**

| Event | Data Imputed After Event? | Strategy |
|---------------------------------------------------------------------------------|---------------------------|------------------|
| Death | Yes | Hypothetical |
| Loss to follow-up | Yes | Hypothetical |
| Withdrawal | Yes | Hypothetical |
| Treatment discontinuation | No* | Treatment policy |
| Treatment crossover | No* | Treatment policy |
| Receipt of treatment for a condition other than amblyopia | No* | Treatment policy |
| Receipt of alternative treatment for amblyopia (i.e., not allowed per protocol) | No* | Treatment policy |

168 * Observed data will be used for analyses.

169 **9. Missing Data**

170 In general, the procedure for handling missing data is outlined in each section. Where not
 171 otherwise specified, missing data will be excluded, and only complete cases will be analyzed.

172 **10. Intervention Adherence**

173 The number of participants stopping study treatment along with reasons for stopping treatment
 174 will be tabulated for each group.

175 **10.1. Primary Outcome Follow-up (Randomization to 26-week Visit)**

176 At 13 weeks and 26 weeks, the investigator will assess participant adherence to the assigned
 177 treatment. For each participant randomized to LUMINOPIA, the number of movie-watching
 178 hours will be categorized according to percentage of prescribed treatment time as 75-100%, 50-
 179 75%, or <50%. PATCHING calendar data will not be analyzed other than a subjective
 180 assessment by the investigator of adherence at 13, and 26-weeks as Excellent, Good, Fair, or
 181 Poor after review of calendar and interview with parent. The tabulation of data related to
 182 treatment adherence is intended for exploratory purposes only, and therefore formal comparisons
 183 between treatment groups will not be performed.

184 **10.2. Post-primary Outcome Follow-up (39-week Visit and 52-week Visit)**

185 At 39-weeks and 52-weeks for PATCHING randomized participants that choose to receive
 186 Luminopia, the total amount of time watching movies for each treatment period will be

187 categorized by percentage and tabulated as described in section 10.1. However, no formal
188 analyses will be conducted using adherence data.

189 **11. Protocol Adherence and Retention**

190 Protocol deviations and visit completion rates (excluding participants who die before the end of
191 the visit window) will be tabulated for each treatment group. A CONSORT² flow diagram will
192 be constructed showing the following for each group:

- 193 • Numbers of participants who were randomly assigned, received intended treatment, and
194 were analyzed for the primary outcome
- 195 • Losses and exclusions after randomization, together with reasons

196 The number of participants who were consented but not randomized will also be provided.
197 Reasons for not randomizing will not be systematically collected.

198 **12. Safety Analyses**

199 The cumulative proportions of each of the following adverse events by treatment group will be
200 assessed at the initial study phase (enrollment to 26 weeks) and during the LUMINOPIA
201 treatment phase for those originally randomized to PATCHING (26 weeks to 52 weeks). During
202 the initial study phase, the proportions will be compared statistically with Barnard's
203 Unconditional Exact Test considering the number of participants per group as fixed. As type II
204 error (false negative) is more of a concern than type I error (false positive) in safety analyses, we
205 will use $p \leq 0.05$, without adjustment for multiplicity, to define statistical significance in all safety
206 analyses. It is noted that this study is not powered for safety analyses and that absence of a
207 significant effect cannot be taken as evidence that a difference does not exist. The proportion of
208 adverse events occurring during the LUMINOPIA treatment phase for original PATCHING
209 participants will be tabulated.

- 210 • Worsening of best-corrected fellow-eye distance VA of 0.2 logMAR or more
- 211 • New onset strabismus $>5 \Delta$ by SPCT in participants with no strabismus at baseline
- 212 • Strabismus $>10 \Delta$ by SPCT in participants with strabismus at baseline
- 213 • Parental report of diplopia occurring more than once per week
- 214 • Skin irritation
- 215 • Headache
- 216 • Eyestrain
- 217 • Dizziness
- 218 • Night terrors
- 219 • Eye twitching
- 220 • Facial redness

221 The PEDIG DSMC will review safety data tabulated by treatment group at each of its semi-
222 annual meetings and can request formal statistical comparison of any safety outcome at any time
223 if they have cause for concern.

224 **13. Baseline Descriptive Statistics**

225 Demographic and clinical characteristics at enrollment will be tabulated by randomized
226 treatment group, and summary statistics appropriate to their distributions will be reported.
227 Variables will include participant age, sex, race, ethnicity, prior treatment for amblyopia, ocular
228 alignment, refractive error, binocular function, amblyopic-eye VA, fellow-eye VA, and IOD.

229 **14. Planned Interim Analyses**

230 There is no plan for formal interim analyses. The Data and Safety Monitoring Committee
231 (DSMC) will review tabulated and graphical displays of interim safety data at approximately 6-
232 month intervals and will have the option to recommend stopping the study.

233 **15. Subgroup Analyses**

234 Subgroup analyses will be used to evaluate potential effect modification (interaction) between
235 the randomized treatment and each pre-specified baseline variable. These analyses will be
236 considered exploratory. Missing data will be imputed like the primary analyses except that the
237 subgroup factors of interest, specified below, will be included in the imputation model, which
238 will be stratified by treatment group. Within-subgroup mean differences for the treatment effects
239 with 95% CIs will be estimated for each subgroup by adding an interaction term to the primary
240 analysis models. Results will be presented as forest plots; *p*-values will not be presented.

241 The baseline factors to be evaluated in pre-planned exploratory subgroup analyses include:

- 242 • Amblyopic-eye distance VA
 - 243 ○ Moderate impairment (20/40 to 20/80, 72 to 53 letters)
 - 244 ○ Severe impairment (20/100 to 20/200, 52 to 33 letters)
- 245 • Type of amblyopia
 - 246 ○ Strabismic only
 - 247 ○ Anisometropic only
 - 248 ○ Both strabismic and anisometropic
- 249 • Ocular alignment at near
 - 250 ○ $0 < \text{Heterotropia} \leq 5 \Delta$ by SPCT
 - 251 ○ None by SPCT
- 252 • Prior treatment for amblyopia
 - 253 ○ Yes (prior amblyopia treatment and glasses)
 - 254 ■ ≥ 1 year of treatment
 - 255 ■ < 1 year of treatment
 - 256 ○ No (glasses only)
- 257 • Age
 - 258 ○ 4 to 5 years
 - 259 ○ >5 to 7 years

- 260 • Sex
- 261 ○ Male
- 262 ○ Female
- 263 • Race and Ethnicity
- 264 ○ White and non-Hispanic
- 265 ○ Non-white and/or Hispanic
- 266 • Binocular function at near
- 267 ○ Randot Preschool Stereoacuity (1.6 to 2.9 log seconds of arc)
- 268 ○ Randot Butterfly Stereoacuity (3.3 log seconds of arc)
- 269 ○ Worth 4-Shape Stereoacuity (Fusion)
- 270 ○ Nil stereoacuity

271

272 There are no data to suggest that the treatment effect will vary by sex, race, or ethnicity.

273 However, each of these factors will be evaluated in exploratory subgroup analyses as mandated

274 by National Institutes of Health (NIH) guidelines.

275 If there is insufficient sample size in a given subgroup ($N < 20$), the cut points for continuous

276 measures may be reconfigured to correspond to the observed distribution of values, possibly

277 using the median to determine the cut point.

278 **16. Multiple Comparison/Multiplicity**

279 For the primary outcome, a 95% confidence interval for the treatment group difference will be

280 constructed and used to evaluate the primary hypothesis of non-inferiority and also the

281 possibility of superiority or inferiority (Figure 1).³

282 For the PedEyeQ questionnaire secondary outcomes, the adaptive false discovery rate (FDR)

283 method with two-stage testing will control the FDR at 5% to adjust p -values and confidence

284 intervals for multiplicity.⁴

285 **17. Exploratory analyses**

286 Exploratory analyses will test the null hypothesis of no difference between treatment groups. p -

287 values and CIs will not be adjusted for multiplicity.

288 **17.1. Mean Change in Distance VA at 13 Weeks**

289 Change in amblyopic eye VA from baseline to 13 weeks is a continuous outcome. Analysis,

290 including imputation of missing data, will mirror the primary outcome; the model will be

291 adjusted for baseline amblyopic-eye distance VA and prior treatment for amblyopia.

292 **17.2. Mean Change in Distance VA over 26 weeks (area under the curve)**

293 The change in amblyopic eye distance VA from baseline over 26 weeks (area under the curve)

294 will be calculated for each participant with the trapezoidal rule using the following formula:

295

$$AUC = \sum_{i=1}^n \left(\frac{V_i + V_{i+1}}{2} \times d \right)$$

296 Where V_i is the VA measured at the i^{th} visit, d is the number of days between visits i and $i+1$
297 (based on the target day, not the actual date of completion), and n is the number of outcome
298 visits included in the analysis. This analysis has $n = 3$ as it will include visits at baseline, 13 and
299 26 weeks; note that change in VA is 0 at baseline for all participants. For presentation, the AUC
300 will be divided by the number of days between baseline and the 26-week visit based on the target
301 day (i.e., 182 days) so that the value shown will have units of letters rather than letter-days. The
302 area under the curve can be interpreted as a weighted average of change in VA over 26 weeks
303 with weights proportional to the time between visits.

304 The area under the curve will be calculated after imputation of missing data. The analysis,
305 including imputation of missing data, will mirror the primary outcome; the model will be
306 adjusted for baseline amblyopic-eye distance VA and prior treatment for amblyopia. A boxplot
307 showing AUC for each treatment group over 26 weeks will be constructed.

308 **17.3. Improvement of Amblyopic-eye Distance VA by 2 or More Lines at 13 and 26** 309 **weeks**

310 Improvement of amblyopic-eye distance VA of 2 or more lines (reduction of 0.2 logMAR) at 13
311 and 26 weeks, respectively, will be analyzed as binary outcomes. For each time point, the
312 proportions with improvement ≥ 2 lines and likelihood-ratio 95% CIs for each treatment group
313 will be calculated with logistic regression, adjusting for baseline amblyopic-eye VA and prior
314 treatment for amblyopia.

315 The adjusted risk difference will be calculated using logistic regression with conditional
316 standardization. The delta method will be implemented to construct a 95% CI on the risk
317 difference⁵ and the model-based two-sided p -value will be reported. Missing data will be
318 imputed as described for the primary outcome.

319 **17.4. Binocular Function Change at 13 Weeks and 26 Weeks**

320 The change in binocular function score from enrollment to the 13 and 26-week visit is an ordinal
321 outcome (Table 5). Components of binocularity include results from the following 3 tests:
322 Randot Preschool Stereoacuity (RPS), Random Dot Butterfly, and Preschool Worth 4-Shape
323 (W4S) at near. These tests will create a composite ordinal score of binocular function with 9
324 levels.⁶

325 The difference between LUMINOPIA and PATCHING for the change in binocular function
326 score from baseline to 13 and 26 weeks will be evaluated with the nonparametric Wilcoxon
327 Rank-Sum test. The difference between groups will be estimated using the Hodges-Lehmann
328 estimator with 95% CI. Analyses for binocular function score will be limited to complete case
329 data at each respective outcome visit (13 weeks or 26 weeks).

Table 5. Levels of Binocular Function as Seconds of Arc on Near Stereoacuity Tests

| Seconds of arc | Log ₁₀ seconds of arc | Stereoacuity Test (Measured at Near) | Binocularity Score (Ordinal) |
|----------------|------------------------------------|--------------------------------------|------------------------------|
| 40 | 1.60 | Randot Preschool | 1 |
| 60 | 1.78 | | 2 |
| 100 | 2.00 | | 3 |
| 200 | 2.30 | | 4 |
| 400 | 2.60 | | 5 |
| 800 | 2.90 | | 6 |
| 2000 | 3.30 | Randot Butterfly | 7 |
| n/a (Fusion) | n/a | Worth 4-Shape | 8 |
| n/a (Nil) | n/a | | 9 |
| Missing | Missing (not included in analyses) | | Missing |

331 **17.4.1. Binocular Function Sensitivity Analysis**

332 In a sensitivity analysis, the difference between LUMINOPIA and PATCHING on the change in
333 binocular function score at the 13- and 26-week visits, respectively, will be evaluated with
334 parametric methods to allow adjustment for baseline binocular function score and prior treatment
335 for amblyopia; such methods will also permit imputation of missing data. For this analysis,
336 fusion and nil will be arbitrarily assigned values of 4000 and 8000 arcsec (each double the
337 previous level), respectively. Using ANCOVA, the adjusted mean difference and 95% CI in log₁₀
338 arcsec between the treatment groups will be presented. Missing binocular function data will be
339 imputed using fully conditional specification (FCS) with logistic regression (cumulative logit) in
340 100 imputations.^{7,8} Imputation will be carried out separately for PATCHING and LUMINOPIA.
341 Variables in the imputation models will include prior treatment for amblyopia and binocular
342 function scores at baseline, 13, and 26 weeks. This method of imputation is being used instead of
343 Markov chain Monte Carlo so that the imputed values are consistent with the possible values of
344 log₁₀ arcsec from the binocular function score.

345 **17.5. Resolution of Amblyopia at 13 weeks and 26 weeks**

346 Resolution of amblyopia is defined as ≤ 0 lines IOD and fellow eye VA no worse than 1 line
347 (0.10 logMAR) below baseline. The cumulative probability of amblyopia resolution at 13 and 26
348 weeks will be calculated using Cox proportional hazards regression with direct adjustment.
349 Event times will be grouped based on the target day of the visit; all 13-week visits will have time
350 set to 91 days and all 26-week visits will have time set to 182 days. Ties will be modeled using
351 the exact method. The IOD in VA at baseline will be included as a covariate because the
352 outcome is a function of the IOD; the model will also control for prior treatment of amblyopia.

353 Participants who are lost to follow up will be censored on the day of the last completed visit. For
354 each visit, the rate of resolution (estimated using the survivor function) and 95% CI will be
355 presented for each group along with the difference in rates and 95% CI, and *p*-value (based on a
356 Z test). To aid in interpretation, Kaplan-Meier curves will be plotted and the number of
357 participants at risk will be shown by visit.

358 **17.6. Treatment Impact Questionnaire**

359 The Treatment Impact Questionnaire (TIQ) will be used as a quantitative measure to evaluate
360 child, proxy, and parent opinions regarding the burdens and impact of the randomized treatment
361 at 13 weeks and 26 weeks (as questions for the child – the Child TIQ, for the parent about the
362 child – the Proxy TIQ, and the parent themselves – the Parent TIQ). The Child-TIQ, Proxy-TIQ,
363 and Parent-TIQ will undergo separate factor analysis to determine the number of domains for
364 each TIQ. Each domain will be refined through the evaluation of misfitting items and will then
365 be Rasch scored. Note that because the TIQ is not administered at baseline (because treatment
366 has not been started), there will be no adjustment for baseline score in any analysis.

367 Additional methods to score and analyze the Treatment Impact Questionnaire will be detailed in
368 a separate SAP.

369 **17.7. Post Primary Outcome Follow-up**

370 Participants that were randomized to PATCHING who have 1 line or more (0.10 logMAR) IOD
371 residual amblyopia will be offered dichoptic treatment with LUMINOPIA after 26 weeks. These
372 participants will be followed at 39 weeks and 52 weeks to evaluate safety and efficacy. The same
373 safety, binocular function, and VA outcomes evaluated at 13 and 26 weeks will be evaluated at
374 39 and 52 weeks with 26 weeks considered the baseline visit for the extended follow-up.
375 Analyses will be limited to PATCHING-randomized participants who agree to be treated with
376 LUMINOPIA after completing the 26-week primary outcome visit.

377 **17.7.1. Binocular Function Change at 39 and 52 Weeks**

378 The change in binocular function score at 39 and 52 weeks is an ordinal outcome that will be
379 analyzed as described in section 17.4. The change in binocular function score from 26 weeks to
380 39 and 52 weeks will be summarized using the median and interquartile range and the one-
381 sample Hodges-Lehmann estimator with 95% confidence interval. Analyses for binocular
382 function score will be limited to complete case data at each respective outcome visit (39 weeks
383 or 52 weeks).

384 As a sensitivity analysis, mean change in binocular function score (\log_{10} arcsec) will be
385 estimated using ANCOVA with adjustment for 26-week score. Missing binocular function data
386 will be imputed using fully conditional specification (FCS) with logistic regression (cumulative
387 logit) in 100 imputations. Variables in the imputation models will include binocular function
388 scores at 26, 39, and 52 weeks.

389 **17.7.2. Mean Change in Distance VA at 39 and 52 Weeks**

390 The mean difference and 95% CI of the change in VA from 26 weeks to 39 weeks, and from 26
391 weeks to 52 weeks, will be calculated for these participants using ANCOVA to adjust for 26-
392 week primary outcome VA. Missing data will be imputed using Markov chain Monte Carlo
393 multiple imputation with 100 imputations. Variables in the imputation model will include VA
394 measured at 26, 39, and 52 weeks.

395 **17.7.3. Improvement of Amblyopic-eye Distance VA by 2 or More Lines at 39 and**
396 **52 weeks**

397 Improvement of amblyopic-eye distance VA of 2 or more lines (reduction of 0.2 logMAR) at 39
398 and 52 weeks, respectively, will be analyzed as binary outcomes. For each time point, the
399 proportions with improvement ≥ 2 lines and likelihood-ratio 95% confidence intervals will be
400 calculated in a logistic regression with intercept only. Missing data will be imputed as described
401 in section 17.7.2.

402 **17.7.4. Resolution of Amblyopia at 39 weeks and 52 weeks**

403 For PATCHING-randomized participants treated with LUMINOPIA during the extended follow-
404 up, the cumulative probability of amblyopia resolution (≤ 0 lines IOD and fellow eye VA no
405 worse than 1 line below baseline) at 39 weeks and 52 weeks will be calculated using Cox
406 proportional hazards regression with intercept only. Event times will be grouped based on the
407 target day of the visit; all 39-week visits will have time set to 273 days and all 52-week visits
408 will have time set to 364 days. Ties will be modeled using the exact method. Participants who
409 are lost to follow up will be censored on the day of the last completed visit. For each visit, rate of
410 resolution (estimated using the survivor function) and 95% CI will be presented.

411 **18. Assumptions**

412 All model assumptions including linearity, normality, and homoscedasticity will be verified
413 using graphical methods. If seriously violated, then transformations, robust methods, or
414 nonparametric methods may be used instead.

415 **19. References**

- 416 1. Hatt SR, Leske DA, Castañeda YS, et al. Development of pediatric eye questionnaires for
417 children with eye disease. *Am J Ophthalmol.* 2019;200:201-217.
- 418 2. Schulz KF, Altman DG, Moher D. CONSORT 2010 Statement: updated guidelines for
419 reporting parallel group randomised trials. *BMC Med.* 2010;8:18.
- 420 3. US Food and Drug Administration. *Non-inferiority clinical trials to establish*
421 *effectiveness: Guidance for industry* 2016. FDA-2010-D-0075. Accessed June 28, 2023.
422 [https://www.fda.gov/regulatory-information/search-fda-guidance-documents/non-inferiority-](https://www.fda.gov/regulatory-information/search-fda-guidance-documents/non-inferiority-clinical-trials)
423 [clinical-trials](https://www.fda.gov/regulatory-information/search-fda-guidance-documents/non-inferiority-clinical-trials)
- 424 4. Benjamini Y, Krieger AM, Yekutieli D. Adaptive linear step-up procedures that control
425 the false discovery rate. *Biometrika.* 2006;93(3):491-507.

- 426 5. Austin PC. Absolute risk reductions, relative risks, relative risk reductions, and numbers
427 needed to treat can be obtained from a logistic regression model. *Journal of clinical*
428 *epidemiology*. 2010;63(1):2-6.
- 429 6. Webber AL, Wood JM, Thompson B, E. BE. From suppression to stereoacuity: a
430 composite binocular function score for clinical research. *Ophthalmic Physiol Opt*. 2019;39:53-
431 62.
- 432 7. White IR, Royston P, Wood AM. Multiple imputation using chained equations: Issues
433 and guidance for practice. *Stat Med*. 2011;30(4):377-99.
- 434 8. Berglund PA. Multiple Imputation Using the Fully Conditional Specification Method : A
435 Comparison of SAS ® , Stata , IVEware , and. 2015: